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## On self-repelling walks

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#### Abstract

We investigate the properties of self-repelling walks-otherwise known as 'true' self-avoiding walks-in both one and two dimensions for a range of values of the repulsion parameter $g, 0.2 \leqslant g \leqslant 10.0$. In one dimension we have obtained 24 terms of the generating function of the mean-square end-to-end distance $\left\langle R_{N}^{2}\right\rangle$, while on the two-dimensional square lattice we have obtained 12-15 terms. In one dimension we find the data to be well fitted by $\left\langle R_{N}^{2}\right\rangle=N^{4 / 3}\left(A+B / N^{1 / 3}+C / N+O(1 / N)\right)$ and in two dimensions by $\left\langle R_{N}^{2}\right\rangle=$ $D N|\ln N|^{\alpha}(1+O(\ln |\ln N| / \ln N))$ with $\alpha \approx 0.5$. Estimates of the amplitudes $A$ and $D$ are also obtained.


## 1. Introduction

A new and interesting variation of self-avoiding walks was recently proposed by Amit et al (1983). This variation was named the 'true' self-avoiding walk (TSAW) to distinguish it from the usual self-avoiding walk (SAW). We consider this name quite inappropriate on the semantic grounds that anything called a 'true' self-avoiding walk has a moral obligation to be self avoiding. In fact, tsaws may recross themselves with non-zero probability, unlike saws, which are truly self avoiding. Accordingly, we suggest the name 'self-repelling walk' (SRw) as semantically appropriate.

On a $d$-dimensional hypercubic lattice an SRW may visit any site $i$ adjacent to its current end point with probability $p_{i}$, which depends on the number of times site $i$ has been visited previously, denoted $n_{i}$, and the repulsion $g>0$, through the relation

$$
\begin{equation*}
p_{i}=\exp \left(-g n_{i}\right)\left(\sum_{i=1}^{2 d} \exp \left(-g n_{i}\right)\right)^{-1} \tag{1.1}
\end{equation*}
$$

The probability of a given walk is just the product of the probabilities of the $N$ steps. Like the sAW problem, the sRW problem is also a non-Markovian process.

Note too that all $(2 d)^{N} N$-step pure random walks are possible (with varying probability) so that the chain generating function (CGF) is just $C(x)=1 /(1-2 d x)$, a rather uninteresting quantity. Weighting each sRw by its probability just modifies this quantity to give the even less interesting result for the weighted CGF $C_{\mathrm{w}}(x)=1 /(1-x)$.

In one dimension, the parameter $g$ effectively interpolates between the saw $(g=\infty)$ and the pure random walks ( $g=0$ ). The interpolation is discontinuous as the critical exponents are those appropriate to sRws for finite $g$ and become saw-like only for $g=\infty$. In higher dimensions, however, the $g \rightarrow \infty$ limit does not correspond to saws as is demonstrated explicitly by Amit et al, though of course $g=0$ still corresponds to the pure random walk.

Amit et al also showed that the critical dimensionality of this model is $d_{c}=2$, quite different from the result $d_{c}=4$ which holds for SAws. A renormalisation group (RG) calculation gave

$$
\left\langle R_{N}^{2}\right\rangle \sim A N(\ln N)^{0.4}(1+B \ln |\ln N| / \ln N) \quad \text { for } d=d_{\mathrm{c}}=2
$$

A Monte Carlo study of two-dimensional SRws for $g=0.1,0.3$ and 1.0 supported this form.

In dimensions below the critical dimension Amit et al do not derive critical exponents, but from their results $\beta=-\frac{1}{2} \varepsilon u+\frac{5}{4} u^{2}, u^{*}=\frac{2}{5} \varepsilon$ and $\gamma=\frac{1}{2} u^{*}=\frac{1}{5} \varepsilon$ one may conclude $\left\langle R_{N}^{2}\right\rangle \sim N^{6 \varepsilon / 5}$ or

$$
\left\langle R_{N}^{2}\right\rangle \sim N^{6 / 5}
$$

to first order in $\varepsilon$ for $\varepsilon=d=1$. Subsequently Pietronero (1983) advanced a Flory-type argument that yielded $\nu=\frac{1}{2}$ for $d \geqslant d_{\mathrm{c}}=2$ and $\nu=2 /(2+d)$ for $d<2$ when $\nu$ is the exponent characterising $\left\langle R_{N}^{2}\right\rangle$ through $\left\langle R_{N}^{2}\right\rangle \sim N^{2 \nu}$. Pietronero argued that this result is inapplicable at $d=1, g=\infty$ for which $\nu=1$ holds. Subsequently Obukhov (1984) argued from a small-g expansion that Pietronero's result holds explicitly at $d=1$, giving $\nu=\frac{2}{3}$ for $0<g<\infty$.

More recently, Rammal et al (1984) have carried out a Monte Carlo study on the one-dimensional SRW problem and find $\nu \approx \frac{2}{3}$. They also consider the case $g<0$ and find a saturation effect so that $\lim _{N \rightarrow \infty}\left(R_{N}^{2}\right\rangle=R_{x}^{2}(g)<\infty$ for $g<0$ (this case could be called the self-attracting walk). de Queiroz et al (1984) considered the crossover behaviour as $g \rightarrow 0$ and $g \rightarrow \infty$. From a real-space RG analysis they provide convincing evidence of a value for $\nu$ different from both the SAW and random-walk values, and constant for $0<g<\infty$.

Since the completion of this work we have received a preprint from Stella et al (1984) who generate series expansions for one-dimensional sRws and find $\nu=0.67 \pm 0.04$, in agreement with both Monte Carlo results and our results. Another recent paper by Family and Daoud (1984) provides a Flory theory for sRws, yielding $\nu=2 /(d+2)$ for $d \leqslant d_{\mathrm{c}}=2$, and argues that the SRW models the statistics of a linear polymer in a polydispersed solution.

In this paper we have studied the sRw in both one and two dimensions by generating series expansions for $\left\langle R_{N}^{2}\right\rangle$ for $N \leqslant 24(d=1)$ and $N \leqslant 15(d=2)$ for $0.2 \leqslant g \leqslant 10.0$.

## 2. One-dimensional results

In table 1 we show the coefficients $\left\langle R_{N}^{2}\right\rangle$ for $1 \leqslant N \leqslant 24$ for several values of $g$ in the range $0.2 \leqslant g \leqslant 10.0$. For large $g$, it is clear that $\left\langle R_{N}^{2}\right\rangle \approx N^{2}$, and the maximum value of $N$ we have used ( $N_{\max }=24$ ) is far too small for asymptotic behaviour to be evident. Indeed, the Monte Carlo results of Rammal et al show that, for $g=10$, deviations from $\left\langle R_{N}^{2}\right\rangle \sim N^{2}$ require $N \geqslant 500$ in order to be evident. Thus our analysis will focus on lower values of $g$, which we restrict to $g \leqslant 2.0$. The heuristic arguments of Amit et $a l$, while giving the wrong exponent, do suggest that the correct form for $\left\langle R_{N}^{2}\right\rangle$ should contain at least two terms, one corresponding to the random-walk behaviour ( $\propto N$ ) appropriate to $g=0$ and the other $\left(\propto N^{4 / 3}\right)$ appropriate to $g>0$. Accordingly we write

$$
\begin{equation*}
\left\langle R_{N}^{2}\right\rangle \sim N^{4 / 3}\left(A+B / N^{1 / 3}+C / N+D / N^{4 / 3}+\ldots\right) . \tag{2.1}
\end{equation*}
$$

Table 1. Mean-square end-to-end distances for one-dimensional self-repelling walks.

| $N$ | $g=0.2$ | $g=0.5$ | $g=\ln 2$ | $g=1.0$ | $g=2.0$ | $g=5.0$ | $g=10.0$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 1.00000 | 1.00000 | 1.00000 | 1.00000 | 1.00000 | 1.00000 | 1.00000 |
| 2 | 2.19934 | 2.48984 | 2.66667 | 2.92423 | 3.52319 | 3.97323 | 3.99982 |
| 3 | 3.41854 | 4.09965 | 4.55556 | 5.27557 | 7.20643 | 8.89327 | 8.99927 |
| 4 | 4.77703 | 6.01795 | 6.84444 | 8.16708 | 11.98134 | 15.74737 | 15.99827 |
| 5 | 6.16214 | 8.06862 | 9.35161 | 11.42518 | 24.71501 | 24.52257 | 24.99673 |
| 6 | 7.65766 | 10.36441 | 12.18831 | 15.14939 | 24.39441 | 35.20655 | 35.99455 |
| 7 | 9.17883 | 12.76553 | 15.18974 | 19.15081 | 31.87956 | 47.78643 | 48.99165 |
| 8 | 10.79473 | 15.38219 | 18.47933 | 23.54609 | 40.16359 | 62.25002 | 63.98793 |
| 9 | 12.43374 | 18.08886 | 21.91276 | 28.18309 | 49.13339 | 78.58464 | 80.98330 |
| 10 | 14.15705 | 20.98433 | 25.59365 | 33.15540 | 58.79055 | 96.77827 | 99.97767 |
| 11 | 15.90154 | 23.96355 | 29.40773 | 38.34456 | 69.03724 | 116.81837 | 120.97095 |
| 12 | 17.72231 | 27.10998 | 33.44057 | 43.83054 | 79.88429 | 138.69318 | 143.96306 |
| 13 | 19.56305 | 30.33493 | 37.59538 | 49.51095 | 91.24487 | 162.39027 | 168.95389 |
| 14 | 21.47359 | 33.71136 | 41.94861 | 55.45976 | 103.13694 | 187.89810 | 195.94336 |
| 15 | 23.40328 | 37.16123 | 46.41462 | 61.58710 | 115.48283 | 215.20441 | 224.93138 |
| 16 | 25.39745 | 40.75030 | 51.06207 | 67.95782 | 128.30324 | 244.29781 | 255.91785 |
| 17 | 27.41013 | 44.40859 | 55.81532 | 74.49450 | 141.52945 | 275.16623 | 288.90268 |
| 18 | 29.48289 | 48.19580 | 60.73584 | 81.25504 | 155.18503 | 307.79849 | 323.88580 |
| 19 | 31.57357 | 52.04877 | 65.75603 | 88.16961 | 169.20595 | 342.18262 | 360.86713 |
| 20 | 33.72066 | 56.02198 | 70.93190 | 95.29234 | 183.61909 | 378.30768 | 399.84650 |
| 21 | 35.88514 | 60.05791 | 76.20187 | 102.55888 | 198.36496 | 416.16187 | 440.82391 |
| 22 | 38.10287 | 64.20672 | 81.61778 | 110.01993 | 213.47105 | 455.73441 | 48379922 |
| 23 | 40.33752 | 68.41549 | 87.12286 | 117.61620 | 228.88298 | 497.01358 | 528.77234 |
| 24 | 42.62269 | 72.73083 | 92.76548 | 125.39508 | 244.62860 | 539.98889 | 575.74323 |

By fitting our data to the form (2.1) we find a consistent picture emerges. To be precise, we have solved (2.1) first by truncating at the $\mathrm{O}\left(N^{-1}\right)$ and then at the $\mathrm{O}\left(N^{-4 / 3}\right)$ term. In order to remove the characteristic odd-even oscillation of hypercubic lattice data, we have transformed our series using the transformation (Watts 1974) $z=20 x /(9 x+11)$ where $x$ is the expansion variable in the original generating function $R^{2}(x)=$ $\Sigma_{N>0}\left\langle R_{N}^{2}\right\rangle x^{N}$. This transformation maps the critical point $x_{\mathrm{c}}=1$ to $z_{\mathrm{c}}=1$, but maps the 'antiferromagnetic' critical point $x=-1$ to $z=-10$, far enough away from the radius of convergence that its effect is negligible. After transformation, we denote the transformed quantity by

$$
\begin{equation*}
\left\langle\tilde{R}_{N}^{2}\right\rangle=N^{4 / 3}\left(\tilde{A}+\tilde{B} / N^{1 / 3}+\tilde{C} / N+\tilde{D} / N^{4 / 3}+\ldots\right) \tag{2.2}
\end{equation*}
$$

where $\tilde{A}=A(0.55)^{2 \nu+1}=0.24784 A$.
We find that truncating the series (2.2) after the term $\tilde{C} / N$ gives a very satisfactory fit. The next term, $\tilde{D} / N^{4 / 3}$, in fact slightly improves the quality of the fit, while not significantly changing the leading amplitude $\tilde{A}$. In table 2 we show the results of our analysis for a representative value of $g, g=\ln 2$, where successive triplets of terms from the transformed series $\left\langle\tilde{R}_{N}^{2}\right\rangle,\left\langle\tilde{R}_{N-1}^{2}\right\rangle$ and $\left\langle\tilde{R}_{N-2}^{2}\right\rangle$ are used to find $\tilde{A}, \tilde{B}$ and $\tilde{C}$ in (2.2), and successive quadruplets of terms are used to fit to $\tilde{A}, \tilde{B}, \tilde{C}$ and $\tilde{D}$.

We have extrapolated the sequence of estimates for the leading amplitude $\tilde{A}$ for all values of $g$ used, and find the results for $A$ shown in table 3. Our estimates encompass both sets of results shown in table 2. Because the amplitudes of the

Table 2. Results of fitting the transformed data of table 1 to the form $(2.2)(g=\ln 2)$.

| $\boldsymbol{N}$ | $\tilde{A}$ | $\tilde{B}$ | $\tilde{C}$ | $\tilde{A}$ | $\tilde{B}$ | $\tilde{C}$ | $\tilde{D}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 10 | 0.4322 | -0.4000 | 0.6728 | 0.4931 | -0.6532 | 1.7653 | -1.1339 |
| 11 | 0.4412 | -0.4286 | 0.7155 | 0.5005 | -0.6840 | 1.8984 | -1.2720 |
| 12 | 0.4490 | -0.4543 | 0.7566 | 0.5064 | -0.7094 | 2.0158 | -1.3983 |
| 13 | 0.4559 | -0.4775 | 0.7958 | 0.5111 | -0.7304 | 2.1196 | -1.5135 |
| 14 | 0.4619 | -0.4984 | 0.8334 | 0.5150 | -0.7482 | 2.2129 | -1.6203 |
| 15 | 0.4672 | -0.5174 | 0.8693 | 0.5183 | -0.7636 | 2.2975 | -1.7197 |
| 16 | 0.4719 | -0.5348 | 0.9037 | 0.5210 | -0.7767 | 2.3740 | -1.8117 |
| 17 | 0.4762 | -0.5506 | 0.9366 | 0.5233 | -0.7881 | 2.4427 | -1.8964 |
| 18 | 0.4800 | -0.5652 | 0.9680 | 0.5252 | -0.7978 | 2.5044 | -1.9741 |
| 19 | 0.4834 | -0.5786 | 0.9982 | 0.5269 | -0.8061 | 2.5595 | -2.0448 |
| 20 | 0.4866 | -0.5910 | 1.0270 | 0.5282 | -0.8132 | 2.6082 | -2.1087 |
| 21 | 0.4894 | -0.6024 | 1.0546 | 0.5293 | -0.8192 | 2.6509 | -2.1655 |
| 22 | 0.4920 | -0.6130 | 1.0810 | 0.5302 | -0.8242 | 2.6874 | -2.2152 |
| 23 | 0.4943 | -0.6228 | 1.1063 | 0.5310 | -0.8282 | 2.7180 | -2.2573 |
| 24 | 0.4965 | -0.6319 | 1.1304 | 0.5315 | -0.8313 | 2.7426 | -2.2917 |

Table 3. Leading amplitudes for one-dimensional self-repelling walks.

| $g$ | $A$ |
| :--- | :--- |
| 0.2 | $0.18 \pm 0.03$ |
| 0.5 | $0.42 \pm 0.03$ |
| $\ln 2$ | $0.54 \pm 0.05$ |
| 1.0 | $0.72 \pm 0.06$ |
| 2.0 | $1.3 \pm 0.3$ |

correction terms depend significantly on the set of results used in table 2-that is, on whether we are fitting to three or four parameters in (2.2)—we provide no estimate for these amplitudes. However it is clear from our analysis that $\tilde{B}$ is (algebraically) decreasing as $g$ increases, while the behaviour of $\tilde{C}$ is not monotonic in $g$.

We have also investigated an alternative form to (2.1) and (2.2) in which we include an additional term proportional to $N^{-2 / 3}$. This reflects the possibility that, if there is a 'correction-to-scaling' exponent $\Delta_{1}=\frac{1}{3}$, then there is a second 'correction-to-scaling' exponent of $\Delta_{2}=\frac{2}{3}$. That is, we assume that

$$
\begin{equation*}
\left\langle R_{N}^{2}\right\rangle / N^{4 / 3}=\sum_{k \geqslant 0} a_{k} / N^{k / 3} . \tag{2.3}
\end{equation*}
$$

The apparent convergence of the amplitudes $a_{k}$ is found to be somewhat less rapid and consistent than that obtained from the assumed form (2.1) and (2.2). Under this alternative assumption the leading amplitude is between $1 \%$ and $10 \%$ higher than that given in table 3, for different values of $g$, though in all cases within the error limits quoted.

Thus we conclude that we cannot unequivocally distinguish between (2.1) and (2.3), though the analysis favours (2.1) somewhat. In either event, table 3 contains estimates of the critical amplitudes.

Table 4. Mean square end-to-end distances for two-dimensional self-repelling walks.

| $\boldsymbol{N}$ | $g=0.2$ | $g=0.5$ | $g=\ln 2$ | $g=1.0$ | $g=2.0$ | $g=5.0$ | $g=10.0$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 1.00000 | 1.00000 | 1.00000 | 1.00000 | 1.00000 | 1.00000 | 1.00000 |
| 2 | 2.09494 | 2.21820 | 2.28571 | 2.37538 | 2.55156 | 2.66069 | 2.66663 |
| 3 | 3.19438 | 3.46020 | 3.61224 | 3.82122 | 4.25523 | 4.53964 | 4.55545 |
| 4 | 4.34239 | 4.80191 | 5.06302 | 5.42316 | 6.18863 | 6.71093 | 6.74054 |
| 5 | 5.49456 | 6.16164 | 6.54137 | 7.06589 | 8.18616 | 8.95595 | 8.99970 |
| 6 | 6.67841 | 7.58379 | 8.09674 | 8.80429 | 10.31890 | 11.36773 | 11.42758 |
| 7 | 7.86541 | 9.01775 | 9.66906 | 10.56596 | 12.48086 | 13.79740 | 13.87192 |
| 8 | 9.07574 | 10.49674 | 11.29653 | 12.39552 | 14.73854 | 16.34957 | 16.44062 |
| 9 | 10.28848 | 11.98410 | 12.93577 | 14.24095 | 17.01581 | 18.91114 | 19.01734 |
| 10 | 11.51958 | 13.50633 | 14.61740 | 16.13801 | 19.36341 | 21.56127 | 21.68408 |
| 11 | 12.75261 | 15.03505 | 16.30812 | 18.04737 | 21.72704 | 24.22264 | 24.36130 |
| 12 | 14.00071 | 16.59199 | 18.03293 | 19.99791 | 24.14579 | 26.95235 | 27.10790 |
| 13 | 15.25042 | 18.15427 |  |  |  |  | 29.86422 |
| 14 | 16.51287 | 19.74013 |  |  |  | 32.67856 |  |
| 15 |  | 21.33053 |  |  |  |  |  |

## 3. Two-dimensional results

In table 4 we show our data for $\left\langle R_{N}^{2}\right\rangle$ for the two-dimensional series for $1 \leqslant N \leqslant 15$ and $0.2 \leqslant g \leqslant 10$. The RG calculations of Amit et al suggested asymptotic behaviour of the form

$$
\begin{equation*}
\left\langle R_{N}^{2}\right\rangle \sim N(\ln N)^{a}(D+E \ln |\ln N| / \ln N) \tag{3.1}
\end{equation*}
$$

with $\alpha=0.4$. Subsequently Obukhov and Peliti (1983) disputed this result on the grounds that the calculation of Amit et al assumed that only one coupling constant needed renormalisation in order to remove all infinities in the perturbation theory, while they found that at least two and possibly three coupling constants are involved. As a consequence, with two coupling constants involved, they found $\alpha=1.0$. They also make the point that this value of $\alpha$ will manifest itself earlier-that is for lower $N$ values-the larger the value of $g$. Indeed, for $g=\infty$ Amit et al did extract a small number of walks in their Monte Carlo study which better fitted the form $\left\langle R_{N}^{2}\right\rangle \sim N \ln N$.

We have investigated this disagreement by analysing our series data as discussed below. Note however that the correction term in (3.1) is very slowly varying, and is essentially undetectable by series analysis methods with our range of $N$ values. As $N$ ranges from 10 to 20 , the correction term ranges from $0.362 E$ to $0.366 E$. Thus we expect an 'effective' amplitude of ( $D+0.36 E$ ).

To estimate $\alpha$, we first form the sequence

$$
\begin{equation*}
s_{N}=\left\{\ln \left[\left(\left\langle R_{N}^{2}\right\rangle / N\right) /\left(\left\langle R_{N-2}^{2}\right\rangle /(N-2)\right)\right]\right\} / \ln (\ln N / \ln (N-2)) . \tag{3.2}
\end{equation*}
$$

If $\left\langle R_{N}^{2}\right\rangle \sim D N(\ln N)^{\alpha}$, then $s_{N} \sim \alpha$. The sequence $\left\{s_{N}\right\}$ is defined using alternate terms in $\left\langle R_{N}^{2}\right\rangle$ in order to accommodate the oscillation discussed previously, and the sequence $\left\{t_{N}\right\}$ and $\left\{u_{N}\right\}$, defined by

$$
\begin{align*}
& t_{N}=\frac{1}{2}\left[N s_{N}-(N-2) s_{N-2}\right] \\
& u_{N}=\left[N^{2} t_{N}-(N-2)^{2} t_{N-2}\right] /(4 N-4) \tag{3.3}
\end{align*}
$$

Table 5. Direct estimates $\left(s_{N}\right)$, linear ( $t_{N}$ ) and quadratic ( $u_{N}$ ) extrapolants of the confluent logarithmic exponent $\alpha$ for two-dimensional self-repelling walks.

| $\mathrm{g}=0.2$ |  |  |  | $g=0.5$ |  |  | $g=1$ |  |  | $g=5$ |  |  | $g=10$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $N$ | $s_{N}$ | $t_{N}$ | $u_{N}$ | $s_{N}$ | $t_{N}$ | $u_{N}$ | $s_{N}$ | $t_{N}$ | $u_{N}$ | $s_{N}$ | $t_{N}$ | $u_{N}$ | $s_{N}$ | $t_{N}$ | $u_{N}$ |
| 6 | 0.097 | 0.189 |  | 0.201 | 0.374 |  | 0.308 | 0.543 |  | 0.474 | 0.752 |  | 0.477 | 0.756 |  |
| 7 | 0.117 | 0.204 |  | 0.234 | 0.385 |  | 0.347 | 0.535 |  | 0.504 | 0.660 |  | 0.507 | 0.660 |  |
| 8 | 0.128 | 0.219 | 0.258 | 0.251 | 0.401 | 0.436 | 0.365 | 0.537 | 0.529 | 0.509 | 0.613 | 0.434 | 0.511 | 0.611 | 0.425 |
| 9 | 0.142 | 0.228 | 0.266 | 0.272 | 0.407 | 0.441 | 0.388 | 0.533 | 0.530 | 0.527 | 0.605 | 0.521 | 0.528 | 0.604 | 0.517 |
| 10 | 0.150 | 0.239 | 0.274 | 0.284 | 0.416 | 0.443 | 0.399 | 0.515 | 0.531 | 0.525 | 0.592 | 0.555 | 0.527 | 0.590 | 0.533 |
| 11 | 0.161 | 0.245 | 0.279 | 0.299 | 0.419 | 0.444 | 0.414 | 0.531 | 0.528 | 0.536 | 0.580 | 0.529 | 0.537 | 0.579 | 0.528 |
| 12 | 0.167 | 0.252 | 0.283 | 0.308 | 0.426 | 0.447 | 0.422 | 0.533 | 0.530 | 0.536 | 0.590 | 0.583 | 0.537 | 0.589 | 0.587 |
| 13 | 0.175 | 0.257 | 0.287 | 0.319 | 0.427 | 0.448 |  |  |  |  |  |  | 0.544 | 0.577 | 0.574 |
| 14 | 0.181 | 0.263 | 0.291 | 0.325 | 0.432 | 0.450 |  |  |  |  |  |  | 0.544 | 0.586 | 0.576 |
| 15 |  |  |  | 0.334 | 0.433 | 0.450 |  |  |  |  |  |  |  |  |  |

extropolate alternate $s_{N}$ 's against $1 / N$ and $1 / N^{2}$, in the usual manner of the ratio method. These sequences are shown in table 5, and it can be seen that for $g=0.2$ a rapidly increasing sequence of estimates suggests $\alpha>0.3$. For $g=0.5$ the rate of increase has substantially declined, and supports $\alpha \geqslant 0.45$. For $g=1$ the estimates are quite steady around $\alpha \approx 0.53$, while for $g=5$ and $g=10$ the estimates are generally decreasing, suggesting $\alpha \leqslant 0.6$. Thus for all values of $g$ we have used we find consistent evidence of $\alpha \approx \frac{1}{2}$. This is much closer to the result of Amit et al ( $\alpha=0.4$ ) than that of Obukhov and Peliti ( $\alpha=1.0$ ). It can of course be argued that our series are short, with $N_{\text {max }}=15$, and that asymptotic behaviour only manifests itself for larger values of $N$. Such an objection is entirely valid, but not totally convincing. Firstly, it would be surprising if all values of $g$ pointed to the same (erroneous) value of $\alpha$, and it is the case that a value of $\alpha \approx \frac{1}{2}$ is indicated by all the series. Secondly, for the self-avoiding walk at the critical dimension ( $d=d_{c}=4$ ), a similar number of terms has been shown to be adequate (Guttmann 1978) to determine the correct confluent logarithmic exponent.

It seems worthwhile to conduct a thorough Monte Carlo analysis in order to resolve this point. For intermediate values of $g$, that is $g \approx 1$, the Monte Carlo data of Amit et al with step size up to $N \approx 2^{14} \approx 16000$ favours $\alpha=0.4$ over $\alpha=1.0$, so clearly very large values of $N$ indeed will be needed if this value of $\alpha$ is incorrect. (However, Obukhov and Peliti reanalyse the data of Amit et al at $g=10$ and find support for $\alpha=1.0$.)

We have also estimated the effective critical amplitude by forming the sequence $D_{N}=\left\langle R_{N}^{2}\right\rangle / N(\ln N)^{\alpha}$ with $\alpha=0.5$. Extrapolating alternate terms in the sequence $\left\{D_{N}\right\}$ both linearly and quadratically against $1 / N$, as was done to estimate $\alpha$, we find the following values of the effective amplitude: $0.62(g=0.2), 0.81(g=0.5), 0.90$ $(g=\ln 2), 1.04(g=1.0), 1.4(g=5)$ and $1.4(g=10.0)$. These values depend on the value of $\alpha$, and included possibly substantial contributions from slowly varying correction terms, so their values should not be uncritically accepted. The trend of increasing amplitude with increasing $g$ value, approaching a positive asymptote as $g \rightarrow \infty$, is likely to be correct however. From the Monte Carlo plots of Amit et al, which assume $\alpha=0.4$, we find for the amplitudes $0.43(g=0.1), 0.61(g=0.3)$ and $1.05(g=1.0)$ in reasonable agreement with our own results. Note that natural logarithms are used throughout this paper, whereas some workers use $\log _{2}$.

## 4. Conclusion

We find that the mean-square lengths of one-dimensional SRws are well fitted by (2.1) which implies a correction-to-scaling exponent of $\Delta=\frac{1}{3}$ for all $g$. The critical amplitude $A$ is found to be an increasing function of $g$, and we see that $A / g$ is a decreasing function of $g$, within the range of $g$ values used. This is consistent with the 'small-g' approximation of Obukhov (1984) who gives $A \sim g^{2 / 3}$ for small $g$. An alternative form for $\left\langle R_{N}^{2}\right\rangle$ is given by (2.3), and it is found that this does not fit the data quite as well as (2.1), but does not significantly alter the amplitude estimates.

For two-dimensional SRws we find that the mean-square lengths are well fitted by (3.1), with confluent logarithmic exponent $\alpha \approx 0.5$, in reasonable agreement with the earlier work of Amit et al, rather than the later theory of Obukhov and Peliti. The qualitative behaviour of the critical amplitude is the same as that for one-dimensional walks discussed above.

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